The Cost of Being “True to Yourself” for Mixed Selves: Frame Switching Leads to Perceived Inauthenticity and Downstream Social Consequences for Biculturals

Alexandria L. West¹, Amy Muise¹, and Joni Y. Sasaki²

Abstract

A growing population of biculturals—who identify with at least two cultures—often frame switch, adapting their behavior to their shifting cultural contexts. We demonstrate that frame switching biculturals are perceived as inauthentic by majority Americans and consequently seen as less likable, trustworthy, warm, and competent compared to biculturals who do not frame switch or a neutral control (Studies 1–3, N = 763). In Study 2, describing the bicultural’s behavior as authentic despite its inconsistency partly alleviated the negative effects of frame switching. In our preregistered Study 3, majority American women were less romantically interested in and less willing to date a bicultural who frame switched in his dating profiles (mediated by inauthenticity). The way biculturals negotiate their cultures can have social costs and create a barrier to intercultural relations.

Keywords
frame switching, authenticity, bicultural, multicultural, intercultural relations, intergroup relations

Do I contradict myself? Very well then I contradict myself, (I am large, I contain multitudes.)
—Walt Whitman, “Song of Myself”

Inconsistency Signals Inauthenticity

Western cultures emphasize the individual as an autonomous agent, ideally uninfluenced by external forces (Nisbett et al., 2001). Behaviors ought to reflect one’s singular, true self and not change across situations (Cross et al., 2003). People who behave inconsistently are seen as inauthentic (Kashima et al., 2004), and authenticity is upheld as a virtue (Kernis & Goldman, 2006). This is problematic for biculturals because frame switching requires changing the way they behave according to the cultural context, and this inconsistency might undermine their perceived authenticity and have downstream social costs.

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Social Costs of Inauthenticity

One reason biculturals frame switch is to gain acceptance by being mindful of each cultures’ norms and values. Paradoxically, frame switching may undercut biculturals’ acceptance in Western society because switching between cultural frames violates the dominant culture’s expectation of behavioral consistency (English & Chen, 2011). The social consequences of frame switching may be far from trivial as inauthenticity comes with many costs.

At a person-perception level, frame switching may damage general impressions of biculturals as a fallout of being seen as inauthentic. Extant research with majority Americans shows that perceived authenticity strongly relates to impressions of likeability and trustworthiness (Krumhuber et al., 2007). Further, authenticity is related to perceptions of warmth and competence (West et al., 2018), which are considered universal dimensions in impression formation (Cuddy et al., 2008). Thus, we hypothesize that a bicultural’s frame switching will undermine their perceived authenticity, and subsequently, their likeability, trustworthiness, warmth, and competence.

Frame switching may have additional, powerful consequences for biculturals, particularly in romantic relationships. In Western societies, feeling and being perceived as authentic is fundamental to forming and maintaining romantic relationships (Josephs et al., 2019), and perceived inauthenticity can diminish relationship satisfaction, commitment, and support (Lopez & Rice, 2006; Wickham, 2013). Our final study examines the consequences of frame switching on biculturals’ online dating prospects, an impactful real-world context in which concerns about authenticity are heightened (Toma et al., 2008).

In the current research, we test our key prediction that frame switching undermines a bicultural’s perceived authenticity, subsequently damaging general impressions and their romantic relationship prospects in America. All studies’ materials, data, syntax, and the preregistration for Study 3 are available on Open Science Framework (osf.io/4397c/); pretests, power analyses, additional and meta-analyzed results are also in the Online Supplementary Materials (OSMs).

Study 1

We hypothesize that participants will see a bicultural as less authentic if he frame switches than if he does not and that this reduction in perceived authenticity will have downstream consequences such that the bicultural will be seen as less likeable, trustworthy, warm, and competent.

Method

Participants

Majority Americans (N = 150) participated online via Prolific. Power analyses (α = .05) based on the effect size of frame switching (vs. no switching) on authenticity obtained in a pilot study (d = 2.04; West et al., 2018) indicated 99.9% power with N = 150. To be eligible, participants had to be White, U.S. citizens, born and residing in the United States, English as first language, and had parents born in the United States, Canada, or Western Europe excluding Southern Europe (Lalonde et al., 2013; n = 9 excluded). We excluded participants who failed more than one of four attention checks (recall the bicultural’s name and cultures, n = 8) or indicated that they did not complete the study honestly and attentively (self-report item, n = 0). Final sample N = 133 (57 females, M_age = 34.38, SD_age = 13.46).

Procedure

Following informed consent, participants were randomly assigned to read one of three vignettes: (1) Switching (n = 44), the bicultural’s behavior differed depending on which cultural group he is with; (2) No Switching (n = 46), the bicultural’s behavior was the same regardless of which cultural group he was with; or (3) Neutral (n = 43), only background information and none on how a bicultural behaved with his cultural groups. After reading the vignette and answering attention checks, participants reported their impressions of the bicultural’s authenticity and provided their impressions of their likeability, trustworthiness, warmth, and competence. Finally, participants completed demographics before debriefing.

Materials

Bicultural vignettes. Participants read vignettes featuring Miguel Wong, a U.S.-born Mexican Chinese bicultural American (West et al., 2018). We selected two minority cultures as the focus of switching to avoid any confounding effects of in-group/out-group biases (e.g., concerns about disloyalty). Both cultures represented minority out-groups for participants, which isolates the effects of frame switching from group biases that may occur if the bicultural was switching between his majority Americans and a minority culture. Vignettes began with the same description of Miguel as an American graduate student who identifies equally with his father’s Chinese culture and his mother’s Mexican culture. The next part of the vignette differed by condition.

The Switching condition described, “Miguel behaves differently depending on which cultural group he is with, so his behavior is more typically Chinese when he is with Chinese people, and more typically Mexican when he is with Mexicans” and then provided examples of how his behavior changes with each culture.

The No Switching condition described, “Miguel doesn’t tend to behave any differently depending on which cultural group he is with, so his behavior is largely the same regardless of whether he is with Chinese people or Mexicans” and provided examples of how he behaves with each culture.

The Neutral condition vignette did not provide any additional information.

Pretesting ensured that the descriptions of Miguel’s specific behaviors did not differ in desirability by condition.
**Table 1. Descriptive Statistics for Study 1.**

<table>
<thead>
<tr>
<th>Outcome Variable</th>
<th>Switching M [95% CI]</th>
<th>No Switching M [95% CI]</th>
<th>Neutral M [95% CI]</th>
</tr>
</thead>
<tbody>
<tr>
<td>Likeability</td>
<td>5.23 [4.98, 5.47]</td>
<td>5.58 [5.34, 5.81]</td>
<td>5.71 [5.46, 5.95]</td>
</tr>
<tr>
<td>Trustworthiness</td>
<td>5.02 [4.70, 5.35]</td>
<td>5.76 [5.44, 6.08]</td>
<td>5.49 [5.16, 5.82]</td>
</tr>
</tbody>
</table>

**Authenticity.** We adapted a four-item measure of subjective authenticity (English & Chen, 2011; α = .94) to assess a target’s perceived authenticity rather than one’s own authenticity, for example, “Miguel is being himself with others” (1 = strongly disagree to 7 = strongly agree).

**General impressions**

**Likeability.** Participants responded to nine items gauging how likeable they found the bicultural (Cila & Lalonde, 2019; α = .88), for example, “Miguel seems like a really nice guy” (1 = strongly disagree to 7 = strongly agree).

**Trustworthiness.** A single item asked, “Overall, I think Miguel is a trustworthy person” (1 = strongly disagree to 7 = strongly agree).

**Warmth and competence.** Participants also rated two fundamental trait dimensions: warmth (six items; α = .87) and competence (seven items; α = .84; Cuddy et al., 2007) on 5-point scales (1 = not at all to 5 = extremely).

**Results**

See Table 1 for descriptive statistics.

A one-way analysis of variance (ANOVA) revealed that authenticity ratings differed significantly across conditions, F(2, 130) = 82.11, p < .001, ηp² = .56. Consistent with our primary hypothesis, participants saw Miguel as less authentic when he frame switched compared to when he actively did not frame switch, t(130) = 12.38, p < .001, d = 2.17, and to when no information was given about his behavior, t(130) = 9.08, p < .001, d = 1.59.

One-way ANOVAs on likeability, trustworthiness, warmth, and competence showed significant effects of condition, F$s$(2, 130) > 4.18, ps < .02, ηp² s > .06 (Table 2—total effects). Across all measures, participants in the Switching condition formed less favorable impressions of Miguel compared to those in the No Switching condition, likeable t(130) = 2.07, p = .04, d = 0.36; trustworthy t(130) = 3.20, p = .002, d = 0.56; warm t(129) = 3.00, p = .003, d = 0.53; competent t(129) = 3.56, p = .001, d = 0.63, and compared to those in the Neutral condition, likeable t(130) = 2.79, p = .006, d = 0.49; trustworthy t(130) = 1.98, p = .05, d = 0.35; warm t(129) = 2.51, p = .01, d = 0.44; competent t(129) = 2.64, p = .009, d = 0.47.

To test whether frame switching negatively affected general impressions by reducing authenticity, we conducted mediation analyses using PROCESS (Version 3) following procedures for multicategorical independent variables (Hayes & Preacher, 2014). Conditions were coded into two orthogonal contrasts: Switching versus No Switching and Switching versus Neutral. Supporting our prediction, confidence intervals for all indirect effects were below zero (Table 2—indirect effects), demonstrating that frame switching significantly decreased evaluations on all traits by diminishing Miguel’s perceived authenticity.

**Study 2**

Study 1 demonstrated that majority Americans saw a frame switching bicultural as less authentic compared to when he actively did not frame switch and when no information was given about his behavior. This loss of perceived authenticity consequently damaged general impressions of the frame switching bicultural. However, statistical mediation in cross-sectional designs is limited to only testing a correlation between the mediator and outcome (Spencer et al., 2005); thus, our next study manipulates the mediator to establish a causal chain between frame switching to authenticity to general impressions. If the consequences of frame switching are truly due to perceived inauthenticity, then assuring participants that a frame switching bicultural is still being authentic should mitigate the harsher impressions found in Study 1. We predicted that majority Americans would form more favorable impressions of a frame switching bicultural when told that he is behaving authentically with each culture compared to when his authenticity is not affirmed.

**Method**

**Participants**

Majority Americans (N = 435) participated online via Prolific. Eligibility and exclusion criteria were consistent with Study 1; final sample N = 390. Power analyses based on an initial study (see OSM) indicated that N = 390 provided 80% power (α = .05) to detect the smallest observed effect—authentic switching versus switching on competence, d = 0.29.

**Procedure**

Overall, the design and procedure followed Study 1. The major difference was adding a new Authentic Switching condition (n = 129) that was based on the previous Switching condition vignette but included an additional paragraph affirming the bicultural’s authenticity. This study also included the same Switching (n = 132) and No Switching (n = 129) conditions from Study 1, allowing us to test whether the previous effects replicated along with the current hypothesis. Thus, there were three randomly assigned conditions: Switching, Authentic Switching, and No Switching. After reading one of the vignettes, participants rated the bicultural’s likeability, trustworthiness, warmth, and competence. They also rated the...
bicultural’s authenticity as a manipulation check before completing demographics and debriefing.

Materials

Bi-cultural vignettes. The Switching and No Switching vignettes were identical to those in Study 1. The new Authentic Switching vignette provided the same content as the Switching vignette, followed by information affirming the bicultural’s authenticity:

Miguel is not trying to pretend or misrepresent himself when he is with either cultural group, and he has no intention to deceive or manipulate others through his behaviour. Rather, Miguel’s behavior with each cultural group reflects different sides of himself that are both equally a part of who he truly is.

Pretesting these vignettes confirmed the effectiveness of the manipulation—Miguel was deemed more authentic in the Authentic Switching (vs. Switching) condition.

General impressions. Measures of likeability (α = .91), warmth (α = .89), and competence (α = .86) were the same as in Study 1. To improve our assessment of trustworthiness beyond a single item, we adapted a three-item measure (Fletcher et al., 2000; α = .93). All response scales ranged from 1 (strongly disagree) to 7 (strongly agree).

Authenticity. The authenticity measure from Study 1 provided a manipulation check; results ensured that the authenticity manipulation in the Authentic Switching condition was successful.

Results

See Table 3 for descriptive statistics.

One-way ANOVAs revealed significant differences between the three conditions on all four general impressions: likability, $F(2, 385) = 5.28, p = .005, \eta^2_p = .03$; trustworthiness, $F(2, 387) = 8.42, p < .001, \eta^2_p = .04$; warmth, $F(2, 385) = 6.70, p = .001, \eta^2_p = .03$; and competence, $F(2, 385) = 8.37, p < .001, \eta^2_p = .04$. Negative effects of the Switching (vs. No Switching) condition also replicated on all impressions, $t(385–387) > 3.17, ps < .002, ds > 0.32$. Assuring participants of Miguel’s authenticity when frame switching (i.e., Authentic Switching vs. Switching) partially mitigated the negative consequences of frame switching, Miguel was judged less harshly in terms of likeability, $t(385) = 2.19, p = .03, d = 0.22$, and warmth, $t(385) = 2.65, p = .008, d = 0.27$, but not trustworthiness, $t(387) = 1.00, p = .32, d = 0.10$, or competence, $t(385) = 0.98, p = .33, d = 0.10$. Further, affirming Miguel’s authenticity when frame switching partially nullified the benefits of actively not frame switching (i.e., Authentic Switching vs. No Switching), as his perceived likeability and warmth did not differ significantly between these two conditions: likeability, $t(385) = 0.97, p = .34, d = 0.10$; warmth, $t(385) = 0.85, p = .40, d = 0.09$. However, actively not frame switching still produced advantages over authenticity frame switching (i.e., No Switching vs. Authentic Switching) for Miguel’s perceived trustworthiness, $t(387) = 2.94, p = .004, d = 0.30$, and competence, $t(385) = 2.95, p = .003, d = 0.30$. Thus, affirming the bicultural’s authenticity countered some, but not all, of the costs from frame switching as well as the benefits from actively not frame switching.

Study 3

Next, we raise the stakes on the social consequences by examining how frame switching negatively impacts biculturals’ romantic relationship prospects. We also address two limitations of the prior studies. First, Studies 1–2 used vignettes explicitly describing the bicultural’s frame switching and so may have had high demand characteristics—participants may have felt expected to react negatively to the bicultural’s inconsistency. Although we would argue that the demand characteristics are likely outweighed by the social desirability of not appearing prejudiced (McConahay et al., 1981), we improve our manipulation in Study 3 to be less explicit by using online

Table 2. Total and Indirect Effects of Frame Switching (vs. No Switching and vs. Control) for Study 1.

<table>
<thead>
<tr>
<th>Outcome Variable</th>
<th>Total Effect of Condition</th>
<th>Switch vs. No Switch via Authenticity</th>
<th>Switch vs. Neutral via Authenticity</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>$F$</td>
<td>$p$</td>
<td>$\eta^2_p$</td>
</tr>
<tr>
<td>Likeability</td>
<td>4.18</td>
<td>.02</td>
<td>.06</td>
</tr>
<tr>
<td>Trust</td>
<td>5.20</td>
<td>.006</td>
<td>.07</td>
</tr>
<tr>
<td>Warmth</td>
<td>5.18</td>
<td>.007</td>
<td>.07</td>
</tr>
<tr>
<td>Competence</td>
<td>6.84</td>
<td>.002</td>
<td>.10</td>
</tr>
</tbody>
</table>

Note. For total effects, $df = 2, df_p = 130$. For indirect effects, nonitalicized coefficients refer to the unstandardized indirect effects, and italicized coefficients below refer to the partially standardized indirect effects. The 95% bias-corrected bootstrap confidence intervals (10,000 samples) that do not contain zero indicate a statistically significant effect. All indirect effects above are significant.
Table 3. Descriptive Statistics for Study 2.

<table>
<thead>
<tr>
<th>Outcome Variable</th>
<th>Authentic Switching M [95% CI]</th>
<th>Switching M [95% CI]</th>
<th>No Switching M [95% CI]</th>
</tr>
</thead>
<tbody>
<tr>
<td>Likeability</td>
<td>5.56 [5.40, 5.72]</td>
<td>5.31 [5.16, 5.47]</td>
<td>5.67 [5.51, 5.82]</td>
</tr>
<tr>
<td>Trustworthiness</td>
<td>5.12 [4.95, 5.30]</td>
<td>5.00 [4.82, 5.17]</td>
<td>5.49 [5.32, 5.67]</td>
</tr>
</tbody>
</table>

...demonstrating his (in)consistency. The No Switching condition presented the same content in each profile with slight variations in how statements were worded to isolate the effects of actively not frame switching from a more mundane form of consistency (i.e., exactly duplicating content).

Participants opened website links to pdfs of Miguel’s profile(s) and were instructed to review them carefully. After freely perusing the profiles, we directed participants’ focus to key aspects with attention checks about the profile photo and content. Participants then rated Miguel on authenticity, general impressions from prior studies, and new dating-relevant impressions. Further, we assessed hypothetical dating intentions toward Miguel. Finally, participants were informed there were no other profiles currently available to rate and provided demographics before debriefing.

Materials

Bicultural dating profiles. All participants saw either one (Neutral) or two dating profiles (Switching or No Switching). All profiles contained the same basic information about Miguel’s demographics, lifestyle, and cultural background. His profile photos (Figure 1) and subtle aspects of the profile content varied between conditions.

In the Switching condition, Miguel had profiles on two real cultural-niche dating websites: MexicanCupid.com and ChinaLoveCupid.com. His MexicanCupid profile photo showed him wearing a shirt with a Calavera (Day-of-the-Dead skull), and the profile content emphasized his interest in more Mexican-associated foods, hobbies/sports, and travel. In contrast, his ChinaLoveCupid.com profile photo showed him wearing a shirt with a Chinese dragon, and the profile content emphasized his interest in more Chinese-associated foods, hobbies/sports, and travel. Importantly, nothing stated in either profile was mutually exclusive—for instance, saying he visited Mexico City in one profile does not contradict the trip to Beijing described in his other profile.

In the No Switching condition, Miguel also had two profiles on the same two cultural-niche websites. In both his MexicanCupid and ChinaLoveCupid profile photos, he was wearing a blank shirt, and the content described his interest in international foods, exercise and sports in general, and a trip to Sydney. Again, the intention here was to demonstrate Miguel’s active nonswitching with culturally neutral content.

In the Neutral condition, Miguel had just one profile on the fabricated, culturally neutral LoveCupid.com, which we created by covering elements of the ChinaLoveCupid.com layout. His photo showed him wearing the same blank shirt, and the profile content was the same as the No Switching condition.

Pretests ensured that participants noticed Miguel’s frame switching between profiles in the Switching condition and made the intended cultural associations (e.g., recognized highlighting of Mexican/Chinese culture) and did not see Miguel as more or less American in the Switching versus No Switching profiles.
Authenticity. Measured the same as previous (α = .93). Two additional exploratory mediators, deceptiveness and manipulativeness, assessed malicious forms of inauthenticity.

Dating-relevant impressions. Impressions of Miguel as a potential dating partner were assessed using a four-item measure of Interpersonal Attraction and Intentions to Meet (Alves, 2018; α = .94), for example, “How much would you like to meet Miguel?” (1 = not at all to 9 = extremely). We also created two new items to assess how attractive (physically and more broadly) participants found Miguel to be (α = .84) and another two items to assess how interested participants were in Miguel as a dating partner (α = .95), for example, “Miguel seems like someone I would be open to dating” (1 = strongly disagree to 7 = strongly agree). Participants also reported how likely they would be to recommend Miguel as a dating partner to a friend using an existing dating endorsement item (1 = strongly disagree to 5 = strongly agree; Rycyna et al., 2009).

Dating intentions. Next, participants indicated how likely they would be to engage in three dating behaviors with Miguel (α = .91). Imagining they had come across Miguel’s profile(s) outside of this study, participants reported their willingness to (1) send Miguel a message, (2) respond to a message from Miguel, and (3) go on a date with Miguel (1 = strongly disagree to 7 = strongly agree).

General impressions. Participants also evaluated Miguel’s likeability, trustworthiness, warmth, and competence using four single-item measures (1 = strongly disagree to 7 = strongly agree), for example, “Overall, I think Miguel is a likeable person.”

Results

Effects on authenticity. One-way ANOVAs revealed significant differences between conditions on ratings of authenticity, $F(2, 237) = 56.21, p < .001, \eta_p^2 = .32$. Participants who witnessed Miguel’s frame switching saw him as less authentic compared to both control conditions (No Switching and Neutral): Switching versus No Switching, $t(237) = 9.68, p < .001, d = 1.26$, Switching versus Neutral, $t(237) = 8.55, p < .001, d = 1.11$. Miguel was not seen as any more or less authentic when he actively did not frame switch (No Switching) compared to when no information about his behavior was given (Neutral), $t(237) = 1.15, p = .25, d = 0.15$. Thus, frame switching had strong negative effects on authenticity, the proposed mediator.1

Consequences for dating-relevant impressions. ANOVA results indicated significant differences between conditions on each of the dating-relevant impressions, $F$s$(2, 237) < 8.48, ps < .001, \eta_p^2$s > .07 (Table 5—total effects). When Miguel frame switched instead of actively not switching or when only one noncultural-niche profile was presented, majority American women formed less favorable dating-relevant impressions. Miguel’s frame switching reduced participants’ Interpersonal Attraction and Intentions to Meet, $t$s$(237) > 2.90, ps < .004, ds > 0.38$; their attraction to him physically and more broadly, $t$s$(237) > 2.90, ps < .004, ds > 0.38$; their interest in him as a dating partner, $t$s$(237) > 2.70, ps < .007, ds > 0.35$; and their endorsement of him as a dating partner, $t$s$(237) > 5.61, ps < .001, ds > 0.07$. To test the role of authenticity as mediating...
Table 4. Descriptive Statistics for Study 3.

<table>
<thead>
<tr>
<th>Outcome Variable</th>
<th>Switching M [95% CI]</th>
<th>No Switching M [95% CI]</th>
<th>Neutral M [95% CI]</th>
</tr>
</thead>
<tbody>
<tr>
<td>Authenticity</td>
<td>4.07 [3.84, 4.30]</td>
<td>5.69 [5.45, 5.92]</td>
<td>5.49 [5.26, 5.73]</td>
</tr>
<tr>
<td>Interpersonal attraction and intentions to meet</td>
<td>4.58 [4.19, 4.97]</td>
<td>5.94 [5.54, 6.33]</td>
<td>5.62 [5.23, 6.02]</td>
</tr>
<tr>
<td>Dating interest</td>
<td>3.33 [2.98, 3.68]</td>
<td>4.35 [4.00, 4.71]</td>
<td>4.01 [3.66, 4.37]</td>
</tr>
<tr>
<td>Dating endorsements</td>
<td>2.80 [2.58, 3.03]</td>
<td>3.82 [3.60, 4.05]</td>
<td>3.71 [3.49, 3.94]</td>
</tr>
<tr>
<td>Interpersonal attraction and intentions to meet</td>
<td>12.66 &lt;.001 .10</td>
<td>-1.47 (24)</td>
<td>-1.80 (.12)</td>
</tr>
<tr>
<td>Attractiveness</td>
<td>11.85 &lt;.001 .09</td>
<td>-0.76 (17)</td>
<td>-1.11 (.11)</td>
</tr>
<tr>
<td>Dating interest</td>
<td>8.48 &lt;.001 .07</td>
<td>-0.80 (20)</td>
<td>-1.32 (.13)</td>
</tr>
<tr>
<td>Dating endorsements</td>
<td>23.84 &lt;.001 .17</td>
<td>-0.60 (14)</td>
<td>-0.88 (.10)</td>
</tr>
<tr>
<td>Dating intentions</td>
<td>7.77 &lt;.001 .06</td>
<td>-0.53 (12)</td>
<td>-0.77 (.12)</td>
</tr>
<tr>
<td>Likeability</td>
<td>24.71 &lt;.001 .17</td>
<td>-0.91 (16)</td>
<td>-1.26 (.16)</td>
</tr>
<tr>
<td>Trustworthiness</td>
<td>37.90 &lt;.001 .24</td>
<td>-0.93 (11)</td>
<td>-1.16 (.11)</td>
</tr>
<tr>
<td>Warmth</td>
<td>11.41 &lt;.001 .09</td>
<td>-0.82 (16)</td>
<td>-1.16 (.11)</td>
</tr>
<tr>
<td>Competence</td>
<td>11.84 &lt;.001 .09</td>
<td>-0.78 (12)</td>
<td>-1.02 (.10)</td>
</tr>
</tbody>
</table>

Note. For total effects, \(df_1 = 2, df_2 = 237\). For indirect effects, nonitalicized coefficients refer to the unstandardized indirect effects, and italicized coefficients below refer to the partially standardized effects. The 95% bias-corrected bootstrap confidence intervals (10,000 samples) that do not contain zero indicate a statistically significant effect. All indirect effects above are significant.

These negative effects, simple mediation models were constructed in line with the analyses described in Study 1. Supporting our hypothesis, confidence intervals for all of the indirect effects were below zero, showing that frame switching significantly diminished majority Americans’ dating-relevant impressions (vs. No Switching and vs. Neutral) because they saw Miguel as less authentic (Table 5—indirect effects). These results show that frame switching in a dating context can make majority Americans feel that a bicultural is being less authentic and, in turn, a less appealing potential romantic partner.

Consequences for dating intentions. The strength of participants’ intentions to communicate with and date Miguel significantly varied between conditions, \(F(2, 237) = 7.77, p = .001, \eta^2_p = .06\) (Table 5—total effects). Participants felt that they would be less likely to send or respond to a message or go on a date with Miguel when he frame switched compared to when he actively did not frame switch and compared to neutral control, \(ts(237) > 2.91, ps < .004, ds > 0.38\). Further, simple mediation results revealed that frame switching reduced participants’ dating intentions (vs. No Switching and vs. Neutral) because Miguel’s frame switching undermined his perceived authenticity (Table 5—indirect effects). Thus, majority Americans were not only less impressed with Miguel as a potential partner when he frame switched but felt they would also be less likely to actually engage with him romantically if they had found these dating profiles on their own in the real world.

Consequences for general impressions. Finally, the results show a significant effect of condition on each of the general
impressions, $F(2, 237) > 11.41, ps < .001, \eta^2_p s > .09$ (Table 5—total effects). Specifically, frame switching cost Miguel in terms of his likeability, trustworthiness, warmth, and competence compared to when he did not frame switch and compared to neutral control, $t(234–237) > 2.77, ps < .006, ds > 0.36$. Mediation analyses confirmed that these consequences of frame switching on general impressions, compared to No Switching and to Neutral, are all mediated by a loss of Miguel’s perceived authenticity when he frame switches (Table 5—indirect effects). These findings directly replicate the second pretest’s results and conceptually replicate each of the earlier studies’ results, adding strong evidence that majority Americans dislike frame switching because they infer that inauthenticity drives the bicultural’s behavior.

General Discussion

Biculturals frame switch as a way to navigate their complex cultural worlds. Across four experiments, however, we demonstrate that frame switching is perceived as inauthentic and, in turn, has social costs. In all studies, the hit to authenticity led to worse impressions of a bicultural’s likeability, trustworthiness, warmth, and competence. Affirming the bicultural’s authenticity in Study 2 partly mitigated frame switching’s negative effects on likeability and warmth, but not trustworthiness or competence. Study 3 targeted the impactful arena of romantic relationships, demonstrating that frame switching in dating profiles diminishes bicultural’s perceived authenticity and reduces their chances of dating success with majority Americans. These results illustrate how frame switching creates a paradox for biculturals living in Western cultures: It allows them to fit in with their cultural groups, but it can backfire when behaving inconsistently violates perceivers’ expectations and values. That is, frame switching biculturals can incur powerful social penalties to impression formation and romantic relationships.

These findings illuminate a novel barrier to intercultural relations in Western society. Previous research has shown that majority Americans are suspicious of biculturals by default because of their dual identities (Kunst et al., 2018) that are assumed to confuse biculturals about who they truly are (Skinner et al., 2019). While these biases may be at play, our results showed that the negative effect on authenticity and its downstream consequences held when frame switching was compared to a neutral control condition in which participants only knew about the bicultural’s dual cultural identities but did not know about his behavior with his cultural groups. This implies that majority Americans’ reactions were driven by the way the bicultural behaved beyond any biases they may hold against his particular cultures or against his dually identified bicultural status.

Recent studies have uncovered that “passing” behavior, whereby a biracial presents as only one racial identity based on the context, also evokes negative reactions from majority Americans (Albuja et al., 2018). In our studies, we were able to isolate a different source of bias against biculturals—switching between their multiple identities—providing some of the first evidence that biculturals’ overt behavior across cultural contexts affects the way they are seen by others. Cumulatively, the previous and current work work unveil the quagmire that biculturals face in Western society—they are punished by majority members not only when they deny one of their identities but also when they present both identities and adapt themselves to their cultural contexts by frame switching. This raises the question: Is there any socially accepted way to be “true to yourself” for mixed selves?

Limitations and Future Research

These studies have some limitations. We only created one bicultural target used across the studies and so we have not examined how target gender or how other minority cultures might change reactions to frame switching. Because Study 3 participants were heterosexual women, we do not know how men or non-heterosexual people would react to prospective bicultural partners’ frame switching. We anticipate that the shared Western understanding of authenticity and its incompatibility with frame switching would be strong enough to influence most majority Americans’ reactions to biculturals, but future research is needed to uncover potential moderators of frame switching’s negative effects. Additionally, the control conditions in these studies depicted a particular form of “not frame switching,” whereby the bicultural’s behavior was intended to be not directly linked to either culture, rather than aligned with one culture over the other (e.g., always more Chinese, as in assimilation) or uniquely mixed together (i.e., hybridizing; West et al., 2017). Future studies should pit frame switching against these and other cultural negotiation strategies for a more complex understanding of how biculturals’ behavior is perceived.

Notably, Study 2 failed to explicitly affirm the biculturals’ authenticity to mitigate the damage of frame switching on trust and competence, even though Studies 1 and 3 establish authenticity as a statistical mediator. It is possible that our manipulation did not cover aspects of authenticity more relevant to trust and competence, or that other mediators may factor more heavily for these two outcomes. Alternatively, affirming authenticity may have weaker benefits for a frame switching bicultural because Americans may not hold an authentic mixed self in as high regard as they would an authentic singular self that personifies their understanding of authenticity. Of these two downstream consequences, implications for trustworthiness are particularly impactful because trust is regarded as fundamental to harmonious relationships (Rempel et al., 1985). This fits well with the results of Study 3, which examined a romantic relationship domain and also suggests that frame switching may lead to particularly harsh penalties in contexts where trust is important. Future research may investigate the fallout of frame switching for bicultural politicians, job applicants, and those already in intercultural romantic relationships. In contrast, Study 2 successfully restored impressions of likeability and warmth by affirming authenticity—results with
implications for ameliorating intercultural relations. At least for these traits, our results demonstrate that Americans can form favorable impressions of a bicultural despite their frame switching. Due to the limits of cross-sectional mediation, these data are not ideally suited to comparing alternate models (e.g., parallel or sequential mediation between perceived authenticity and other trait impressions). Future longitudinal studies should examine how impressions may change and develop over the course of multiple interactions with a frame switching bicultural to more comprehensively test the role of perceived authenticity over time.

Conclusion
A growing population of biculturals endeavor to be true to their mixed selves. However, the strategies biculturals use to successfully navigate their multiple cultures can have social costs. As many nations become increasingly diverse, it is more important than ever to identify and break down these barriers to intercultural relations so that all people can thrive while being true to themselves.

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Supplemental Material
The supplemental material is available in the online version of the article.

Note
1. Miguel was also rated as more deceptive and manipulative in the Switching condition (vs. No Switching and vs. Neutral). When authenticity, deceptiveness, and manipulativeness were entered simultaneously into parallel mediation models, only authenticity produced unique indirect effects consistently across all outcomes.

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Alexandria L. West is a PhD candidate in psychology at York University. Her research tests how the processes biculturals and intercultural couples use to negotiate their cultures affect them psychologically and socially.

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